

**Advanced Engineering Mathematics**  
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**Lecture – 50**  
**Exponential Distribution**

Hello friends welcome to my lecture on exponential distribution let  $N(t)$  be a poisson process with rate  $\lambda$  let  $x_1$  be the time of 1st arrival then  $p$  be the probability that  $x_1 > t$ .

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Arrival and interarrival Time

Let  $N(t)$  be a Poisson process with rate  $\lambda$ . Let  $X_1$  be the time of first arrival. Then,

$$P(X_1 > t) = P(\text{no arrival in } (0, t)) = P(N(t)=0) = \frac{(\lambda t)^0 e^{-\lambda t}}{0!} = e^{-\lambda t}$$

Hence

$$F_{X_1}(t) = \begin{cases} 1 - e^{-\lambda t} & t > 0 \\ 0 & \text{otherwise.} \end{cases}$$

$P(N(t)=n) = \frac{(\lambda t)^n e^{-\lambda t}}{n!}$   
 $P(N(t)=0) = e^{-\lambda t}$   
 $F_{X_1}(t) = P(X_1 \leq t) = 1 - P(X_1 > t) = 1 - e^{-\lambda t}$

Will be = to probability that there is no arrival in the interval 0 to t because  $x_1$  is the time of the first arrival and  $x_1 > t$  so in the time interval 0 to t there will be no arrival and probability that there is no arrival in 0 to t means probability that  $N(t) = 0$  means we have probability of in a Poisson process probability that  $N(t) = n = \lambda t$  rest to the power  $n$   $e$  to the power  $-\lambda t$  /  $n$  factorial so there is no arrival in 0 to t means probability that  $N(t) = 0$ .

$N(t) = 0$  means  $\lambda t$  to the power 0 which is 1 then  $e$  to the power  $-\lambda t$  / 0 factorial so this will be  $e$  to the power  $-\lambda t$  so this probability of no arrival in 0 to t is actually probability of  $N(t) = 0$  there is no occurrence in the time interval 0 to t so this =  $e$  to the power  $-\lambda t$  hence  $F_{X_1}(t)$  is probability that  $x_1 \leq t$  probability that  $x_1 \leq t$  and this = 1 - probability that  $x_1 > t$  so probability at  $x_1 > t$  is  $e^{-\lambda t}$ .

So, it should be  $1 - e^{-\lambda t}$  so we will get this  $1 - e^{-\lambda t}$  when  $t > 0$  and it will be 0 otherwise

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Therefore  $X_1 \sim \text{Exponential}(\lambda)$ . Let  $X_2$  be the time elapsed between the first and the second arrival (fig)

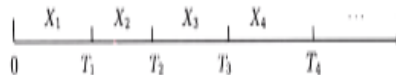


Figure 11.4 - The random variables  $X_1, X_2, \dots$  are called the interarrival times of the counting process  $N(t)$ .

Figure : Fig.1

Now therefore  $x_1$  is an exponential  $\lambda$  distribution and let  $x_2$  be the time elapsed between the 1st and the 2nd arrival so  $x_1$  is the time of the 1st arrival  $x_2$  is the time elapsed between the 1st and the 2nd arrival so then  $x_3$  is the time between the 2nd and 3rd arrival like that.

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Let  $s > 0$  and  $t > 0$ . Note that the two intervals  $(0, s)$  and  $[s, s+t)$  are independent. We can write  $P(X_2 > t | X_1 = s) = P(\text{no arrival in } [s, s+t) | X_1 = s)$

$$\begin{aligned}
 &= P(\text{no arrival in } [s, s+t) | \text{independent increments}) \\
 &= \underline{e^{-\lambda t}} \quad \checkmark \quad P(N(t)=0) = \frac{(\lambda t)^0 e^{-\lambda t}}{0!} = e^{-\lambda t}
 \end{aligned}$$

So, what we have let  $s > 0$  and  $t > 0$  let us note that the 2 intervals  $0$  as intervals  $0$   $s$  and  $s+t$  are independent we know that of the number of occurrences are independent of the non-overlapping intervals if there are 2 over an interval which are non-overlapping number of

occurrences in 1 is independent of the number of occurrences in the other so let us note that 0 as  $s$   $s+t$  are independent therefore probability that  $x_2$  is  $> t$ .

Given that  $X_1 = s$  given that  $X_1 = x_1$  is the time of the 1st arrival  $x_1 = s$  that means  $s$  is the time of the 1st arrival  $x_2$  is  $t$  means there is no arrival in  $s + t$   $X_2$  is the time elapsed between the 1st arrival and the 2nd arrival 1st arrival occurs at  $s$  at time  $s$  so that means and these 2nd the time elapsed between the 1st arrival and 2nd arrival is more than  $t$  it will mean that there is no arrival in the interval  $s$   $s+t$ .

Okay given that  $x_1 = s$  now  $s$   $s+t$  and  $0$   $s$  are independent so probability that no arrival is there in  $s$   $s+t$  interval given that  $x_1 = s$  is system  $s$  probability that there is no arrival in the interval as  $s$   $s+t$  now there is no arrival in  $s$   $s+t$  is same as the probability that there is no arrival in the interval  $0$  to  $t$ , we have already found the probability of no arrival in  $0$  to  $t$  probability of arrival in  $0$  to  $t$  is same as probability of  $n$  arrivals because it follows the same probability distribution.

So, no arrival  $s$   $s+t$  means  $e^{-\lambda t}$  it follows the same probability distribution means no arrival in  $s$   $s+t$  will mean that probability that  $n=0$  so this is the number of occurrences in the time interval  $0$  to  $t$  because of the independence of  $0$  to  $t$   $s$   $s+t$  number of no occurrences in the interval  $s$  to  $s+t$  will be no occurrence in the interval  $0$  to  $t$  so that is probability of  $n=0$ .

We can find so this will be  $\lambda^0 e^{-\lambda t}/0!$  so this is  $e^{-\lambda t}$  so we get the probability of  $x_2 > t$  given that  $x_1 = s$   $e^{-\lambda t}$ .

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We conclude that  $X_2 \sim \text{Exponential}(\lambda)$ , and that  $X_1$  and  $X_2$  are independent. The random variable  $X_1, X_2, \dots$  are called interarrival times of the counting process  $N(t)$ . Similarly, we can argue that all  $X_i$  are independent and  $X_i \sim \text{Exponential}(\lambda)$  for  $i=1,2,3,\dots$

#### Interarrival Times for Poisson Process

If  $N(t)$  is a poisson process with rate  $\lambda$ , then interarrival times  $X_1, X_2, \dots$  are independent and

$X_i \sim \text{Exponential}(\lambda)$  for  $i=1,2,3,\dots$

If  $X$  is exponential with parameter  $\lambda > 0$ , then  $X$  is a memoryless random variable, that is

$$P(X > x + a | X > a) = P(X > x), \text{ for } a, x \geq 0.$$

And thus we can conclude that  $x_2$  is exponential lambda of following exponential lambda distribution thus  $X_1$  and  $X_2$  are we conclude that  $X$  to follow the exponential lambda distribution and that  $X_1$  and  $X_2$  are independent random variables  $X_1$   $X_2$  are called inter arrival times of the counting process and we have seen that  $X_1$  is the time of the 1st arrival  $X_2$  is the time elapsed between the 1st first and 2nd arrival  $x_3$  the time elapsed.

Between 2nd second and 3rd arrival so they are called inter arrival times of the counting process  $N_t$  similarly we can argue that all  $X_i$  are independent and  $X_i$  are follow exponential lambda distribution now inter arrival times for Poisson process if  $N_t$  is the Poisson process with the rate lambda with the rate lambda then inter arrival times  $X_1$   $X_2$  are independent and follow exponential lambda distribution if  $X$  is exponential with parameter lambda  $> 0$ .

Then  $X$  is a memory less random variable so  $x$  is a memory less random variable means the probability that  $x$  is  $> X+a$  given that  $X$  is  $> a$  is same as the probability that  $X > x$  so whatever has happened given that  $X > a$  only depends on the length of  $a$  to  $x + a$  that is  $x$  so probability that  $X$  is  $> x$  later we shall in the lecture we shall discuss in detail about this memory less property of the exponential distribution.

So, we will see that it depends on the length of the interval not on the values of this  $a$  and  $x + a$  we it only depends on the length of the interval.

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### Exponential distribution

Let us define the random variable  $T$  as the waiting time for the first occurrence from time  $t = 0$ , then  $T$  is a continuous random variable since it can take any value in an interval. The distribution function of the random variable  $T$  is obtained by relating it with the discrete random variable  $N(t)$  = number of occurrence in  $(0, t)$ . The event  $\{T > t\}$  is the same as the event  $\{N(t) = 0\}$ . Recall that  $N(t)$  is a discrete random variable with Poisson distribution with parameter  $\lambda t$ .

$$P(N(t) = 0) = (\lambda t)^0$$

Now let us define the random variable  $T$  as the waiting time for the 1st occurrence from time  $t = 0$  so  $t$  is the waiting time for the 1st from time  $t = 0$  then  $T$  is a continuous random variable since it can take any value in an interval the distribution function of the random variable  $T$  is a obtained by relating it with the discrete random variable  $t$  and we know that  $N t$  is the discrete random variable of a continuous variable  $T$  and  $t$  denotes the number of occurrences.

In the interval  $0$  to  $t$  the event  $T > t$  is the same as event  $t = 0$  because  $t$  is the time of the first break even time for the 1st occurrence and if  $T > t$  it is same as and  $t = 0$  there is no occurrence in the interval  $0$  to  $t$  now recall that  $Nt$  is the discrete variable with Poisson distribution with parameter  $\lambda t$  this we have discussed in the previous lecture that  $Nt$  is a discrete  $n$  variable and follows Poisson distribution.

With parameter  $\lambda t$  because the probability here in is given by  $\lambda t$  rest to the power  $n$  and probability that  $Nt = n = \lambda t$  to the power  $n$   $e^{-\lambda t} / n!$  so it follows the Poisson distribution with parameter  $\lambda t$ .

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### Exponential distribution cont...

Hence

$$P(T > t) = P(\{N(t) = 0\}) = \frac{e^{-\lambda t} (\lambda t)^0}{0!} = e^{-\lambda t} \checkmark$$

Thus  $P(T \leq t) = 1 - e^{-\lambda t}, t > 0$ . Hence the distribution function of T is given by

$$F_T(t) = 1 - e^{-\lambda t}, t > 0.$$

This distribution function of T is referred to as the exponential distribution. Thus we conclude that the waiting time for the first occurrence in a Poisson process is distributed in the exponential form.

$$F_T(t) = P(T \leq t) = 1 - e^{-\lambda t}$$

Probability that  $T > t$  is same as we have just now discussed  $t$  is the time left between  $t = 0$  to  $t$  for the 1<sup>st</sup> occurrence so if the time for the 1<sup>st</sup> occurrence is  $t$  and it is  $> t$  that means in the interval  $0$  to  $t$  there is no occurrence so probability that  $N_t = 0$  probability of  $N_t = 0$  is  $e$  to the power  $-\lambda t$   $\lambda t$  to the power  $0/0$  factorial so it is  $e$  to the power  $-\lambda t$  thus probability of  $T < = t$  probability that  $T < = t$  is  $1 -$  probability that  $T > t$ .

That is it is  $1 - e^{-\lambda t}$  to the power  $-\lambda t$   $t > 0$  hence the distribution function of T is given by because  $F_T(t)$  denotes the distribution function of  $t$ , so it is given by probability that  $T \leq t$  and we have seen  $t$  probability  $T \leq t$  is  $1 - e^{-\lambda t}$  so the distribution function of T is given by  $F_T(t)$  and it is  $1 - e^{-\lambda t}$  now this distribution function of T is referred to as the exponential distribution.

Thus we can conclude that the waiting time for the 1<sup>st</sup> occurrence in a Poisson process is distributed in the exponential form it is given by  $1 - e^{-\lambda t}$ .

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The probability density function of T

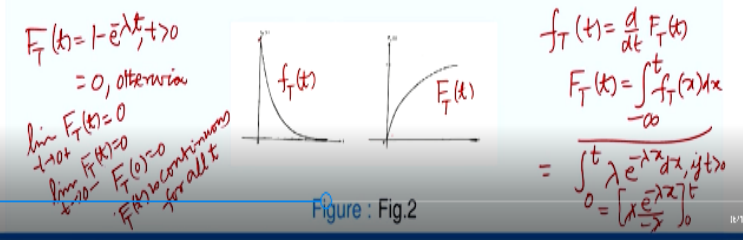
$$f_T(t) = \frac{d}{dt} F_T(t) = \begin{cases} \lambda e^{-\lambda t} & t > 0 \\ 0 & t \leq 0 \end{cases}$$

*Handwritten notes:*  $\lim_{t \rightarrow 0^-} f_T(t) = 0$  &  $\lim_{t \rightarrow 0^+} f_T(t) = \lambda$  since  $\lambda > 0$

This is the probability density function of the exponential distribution with parameter  $\lambda > 0$  for  $t > 0$

$$F_T(t) = \int_{-\infty}^t f_T(x) dx = \int_0^t \lambda e^{-\lambda x} dx = 1 - e^{-\lambda t}$$

The graph of  $f_T(t)$  and  $F_T(t)$  for the exponential distribution are shown below:



Now the probability density function of T we know that the derivative of the distribution function  $F_T(t)$  the cumulative distribution function gives the probability density function so probability density function if we denote by  $f_T(t)$  it is  $d/dt$  of  $F_T(t)$  and  $d/dt$  of  $F_T(t)$  we can find from here.  $F_T(t)$  is  $= 1 - e^{-\lambda t}$  when  $t > 0$  and 0 otherwise so from here  $d/dt$  of  $F_T(t)$  we can determine this  $= \lambda e^{-\lambda t}$  when  $t > 0$ .

And 0 otherwise so as we have just now seen probability density function probability density function is = derivative of the distribution function so a probability density function of the exponential distribution is given by  $f_T(t) = \lambda e^{-\lambda t}$  when  $t > 0$  and 0 otherwise so we get this now this is the probability density function of the exponential distribution with parameter  $\lambda > 0$   $T > 0$   $F_T(t)$  if you integrate  $F_T(t) =$  this.

We earlier also discuss this  $F_T(t)$  when you integrate this equation  $F_T(t)$  is given by the integral/ - infinity to t  $F_T(x) dx$  so the cdf of the cumulative distribution function of t the random variable t is given by integral/ - infinity to t  $F_T(x) dx$ . And here you can easily see that if you put the value of  $F_T(x)$  is = 0 when  $x \leq 0$  so from here we can see this is = integral/ 0 to t  $F_T(x) = \lambda e^{-\lambda x}$  if  $t > 0$  if  $t < 0$  then  $F_T(x)$  will be 0.

So, this cumulative distribution function  $F_T(t) = 0$  and when you integrate this what we get is  $e^{-\lambda x} / -\lambda + C$  and we take 0 to t so this gives you when you can

cancel this lambda. With this lambda and when you substitute the limits you get one minus E to the power - lambda so  $F(t) = 1 - e^{-\lambda t}$  when  $t > 0$  and 0 otherwise so we get back the cumulative distribution function  $F(t) = \int_{-\infty}^t f(x) dx$ .

Now here you can see the graph of this is the graph of density function  $f(t)$  and this is the graph of we have taken  $F(t)$  here we have taken cdf you can see  $F(t)$  is given by  $1 - e^{-\lambda t}$ . When  $t > 0$  and 0 when  $t \leq 0$  this means that  $\lim_{t \rightarrow 0^-} F(t) = 0$  and  $\lim_{t \rightarrow 0^+} F(t) = 0$  when  $t$  goes to 0  $e^{-\lambda t}$  goes to 1 so this goes to 0. So, the value of  $f(t)$  at  $t = 0$  so this 0 point is here.

So, this means that, and lambda is we have lambda to be positive. So, since lambda is  $> 0$  okay  $f(t)$  is discontinuous at  $t = 0$  left and right hand limits are same. You can see the graph also okay when  $t \rightarrow 0^+$  it goes to the value lambda this is lambda here okay and on the left side okay on the negative  $t$  axis  $f(t)$  is everywhere 0. So, it is continuous on the whole  $t$  axis except at  $t = 0$  while the function  $F(t)$  okay.

$F(t)$  is continuous for all values of  $f(t) = 1 - e^{-\lambda t}$  when  $t > 0$  we can find the limit here as  $t \rightarrow 0^+$ . So, when  $t$  goes to  $0^+$   $e^{-\lambda t}$  goes to 1 so  $1 - 1$  is this means 0 we get, and  $\lim_{t \rightarrow 0^-} F(t)$  is also 0 okay and  $t \rightarrow 0$  also so  $F(0)$  is also 0 okay so we have  $F(t)$  is continuous for all  $t$ . Okay this is clear from the graph also okay on the negative  $t$  axis  $F(t)$  takes 0 value for all  $t$ .

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Note that the density function of exponential distribution is discontinuous at  $t = 0$  and continuous elsewhere. The distribution function  $F_T(t)$  is however, continuous everywhere.

### Application

The exponential distribution is used in life testing, reliability, queueing theory and other areas. Life of an electronic fuse or equipment which does not age with time, service time for a customer in a facility, waiting time between two occurrences of a Poisson process are some of the examples of exponential random variable.

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So, the density function of the exponential distribution that is  $f_T(t)$  is discontinuous at  $t=0$  but continuous everywhere else. The distribution function  $F_T(t)$  is however continuous everywhere. Now the exponential distribution is used in life testing, reliability, queueing theory and many other areas. Okay life of an electronic fuse or equipment which does not age with time, service time for a customer in a facility.

Waiting time between two occurrences of a Poisson process are some of the examples of exponential random variable.

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### Example 1

The number of miles that a particular car can run before its battery wears out, is exponentially distributed with an average of 10,000 miles. The owner of car needs to take 5000 miles trip. What is the probability he will be able to complete the trip without having to replace the car battery ?

Ans: 0.604

Mean of the exponential distribution =  $\frac{1}{\lambda} = 10,000$

$\Rightarrow \lambda = \frac{1}{10,000}$

Let  $X$  denote the number of miles the car can run before its battery wears out

then  $P(X > 5000) = \int_{5000}^{\infty} \lambda e^{-\lambda t} dt = \frac{1}{10,000} \int_{5000}^{\infty} e^{-\lambda t} dt = \frac{1}{10,000} \left( \frac{e^{-\lambda t}}{-\lambda} \right)_{5000}^{\infty}$

$= e^{-\frac{5000}{10,000}} = e^{-1/2} = 0.604$

Now let us take some examples suppose we have to find we have to solve this problem, the number of miles that a particular car can run before its battery wears out is exponentially distributed with average of 10000 miles . So, the owner of car needs to take 5000 miles trip . Okay what is the probability he will be able to complete the trip without having to replace the car battery.

Mean  $=1/\lambda$  mean of the exponential distribution is  $=1/\lambda$  so  $1/\lambda=10000$  miles okay which implies  $\lambda=1/10000$  now the owner of car needs to take a 5000 miles trip what is the probability he will be able to complete the trip without having to replace the car battery. So, let us say that X denote the number of miles the car can run before the battery wears out okay.

Okay then we want the probability that X is  $>5000$  if X is more than 5000, he will be able to complete the trip without the battery wearing out. So, this is  $= \int_{5000}^{\infty} \lambda e^{-\lambda t} dt$  okay that is the probability density function of the exponential distribution so  $\lambda$  is  $= 1/10000$   $\int_{5000}^{\infty} e^{-\lambda t} dt$  .

So, this is  $=1/10000 e^{-\lambda t}/-\lambda$  and we have 5000 to infinity okay. Now when T goes to infinity  $\lambda$  is  $>0$   $\lambda$  is  $1/10000$  so  $e^{-\lambda t}$  will go to 0 and therefore this is  $1/10000$  and when it is 5000 okay then, but we will get  $e^{-5000 \lambda}$   $\lambda/\lambda$  okay we will get this okay? Now  $\lambda$  is  $1/10000$  so this  $\lambda$  and this 10000 will cancel okay?

And what we get  $e^{-5000 \lambda}$  so this is  $= e^{-5000 * 1/10000}$ . And this is  $=e^{-1/2}$  okay? you get to the power - 1/2 comes out to be 0.604 so here we will be able to complete probability the trip with the probability of . 6 05 that means there are 6 more than about 60% chance that he will be able to complete the trip without the battery wearing out.

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### Example 2

Suppose the length of a phone call in minutes follows the exponential distribution with parameter  $\lambda = 0.1$ . If someone arrives immediately ahead of you at a public telephone booth, find the probability that, you will have to wait

(a). more than 10 minutes.

(b). between 10 and 20 minutes.

Ans: (a)  $e^{-1}$  (b)  $e^{-1} - e^{-2}$

Let  $X$  denote the waiting time for a phone call

$$P(X > 10) = \int_{10}^{\infty} \lambda e^{-\lambda t} dt$$
$$= \lambda \left[ \frac{e^{-\lambda t}}{-\lambda} \right]_{10}^{\infty} = e^{-10\lambda} = e^{-10(0.1)} = e^{-1}$$
$$P(10 < X < 20) = \int_{10}^{20} \lambda e^{-\lambda t} dt$$
$$= \left[ \frac{\lambda e^{-\lambda t}}{-\lambda} \right]_{10}^{20} = e^{-10\lambda} - e^{-20\lambda} = e^{-1} - e^{-2} \text{ as } \lambda = 0.1$$

Now let us go to our next question suppose the length of phone call in minutes follows the exponential distribution with parameter  $\lambda = 0.1$  if someone arrives immediately ahead of you at a public telephone booth find the probability that you will have to wait more than 10 minutes so let  $X$  denote the waiting time okay for a phone call okay so then we want the probability that you have to wait more than 10 minutes that is actually  $X > 10$ .

So, this will be integral 10 to infinity okay?  $\lambda e^{-\lambda t} dt$  and  $\lambda = 0.1$  okay so we can write  $\lambda$  here and then 10 to infinity  $e^{-\lambda t}$  so integral will be  $e^{-\lambda t} / -\lambda$  and we have this situation this will be = this  $\lambda$  will cancel with this  $\lambda$  and we will have when  $t$  goes to infinity  $e^{-\lambda t}$  will go to 0 and when it is 10 it will be  $e^{-10\lambda}$  is 0.1 okay.

So,  $e^{-10 \cdot 0.1}$  which means  $e^{-1}$  so we will have to wait more than 10 minutes the probability will be  $e^{-1}$  okay now we have to wait between 10 and 20 minutes. So, probability that  $10 < X < 20$  okay so this means we need the probability that means we have to calculate the integral 10 to 20  $\lambda e^{-\lambda t} dt$ . Probability that  $a < X < b$  is given by  $\int_a^b \lambda e^{-\lambda t} dt$  that is the density function.

Okay so we have here  $e^{-\lambda t} / -\lambda$  and we have 10 to 20. Okay so this  $\lambda$  will cancel with this we will get  $e^{-10\lambda} - e^{-20\lambda}$

lambda = 0.1 okay when we shall put the limits okay. So, this is e to the power -1 - e to the power -2 as lambda = 0.1 okay so we get this probability. This is the probability that we have to wait between 10 and 20 minutes to make the call.

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**Example 3**

Students arrive at a local bar and restaurant according to an approximation Poisson process at a mean rate of 30 students per hour. What is the probability that the bouncer has to wait more than three minutes to card the next student?

**Ans:**  $e^{-3/2} = 0.223$ .

*Poisson mean =  $\frac{30}{60} = \frac{1}{2}$  student per minute*

*Let W denote the waiting time then we can expect on an average 2 minutes waiting time between arriving students*

*Mean =  $\frac{1}{\lambda} = 2$*   
 *$\lambda = \frac{1}{2}$*

$$P(W > 3) = \int_3^{\infty} \lambda e^{-\lambda t} dt = \frac{1}{2} \int_3^{\infty} e^{-t/2} dt = \frac{1}{2} \left( \frac{e^{-t/2}}{-1/2} \right) \Big|_3^{\infty} = e^{-3/2} = 0.223$$

Now suppose we have this problem students arrive at a local bar and restaurant according to an approximation Poisson process at the mean rate of 30 students per hour. What is the probability that the bouncer has to wait more than 3 minutes to card the next student? So students arrive at a local bar and turned according to an approximation Poisson process at a mean rate of 30 students per hour meaning that the Poisson mean is 30/60.

Okay 30 students per hour means 1/2 student per minute okay now let us say let W denote the waiting time okay then we can expect on an average 2 minutes waiting time between arriving students . Because we have got 1/2 strength means in 2 minutes 1 student so there will be a we can expect that that will be a time interval of 2 minutes between arriving students. So, we want the probability that the bouncer has to wait more than 3 minutes.

Okay on average the time is 2 minutes between arriving students. But we want the probability that the time I mean waiting time is more than 3 minutes. So, we want the probability that the waiting time is W is more than 3 okay that means integral 3 to infinity lambda e to the power -

lambda t dt. Okay now as i said mean of the exponential distribution is given by lambda okay now on an average what we are getting is the betting time is 2 minutes.

Okay so 1/lambda=2 minutes that means lambda=1/2 okay this lambda is same as the mean of the Poisson process the same lambda lambda =1/2 so 1/2 integral 0 to infinity okay e to the power -t/2 dt because lambda=1/2 okay so we have 1/2 e to the power t/2 -1/2 this will cancel with this and we get what when t goes to infinity it will go to 0 okay, so we get e to the power -3/2 e o the power -3/2 is 0.223.

So, the bouncer will have to wait more than 3 minutes the probability for that is 0.223 okay now the expectation and variance of exponential distribution.

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The slide contains the following text and equations:

Expectation and variance of exponential distribution

The service time  $T$  for a customer follows the exponential distribution with parameter  $\lambda$ .

$$E(T) = \int_{-\infty}^{\infty} t f_T(t) dt = \int_0^{\infty} t \lambda e^{-\lambda t} dt = \int_0^{\infty} \frac{u}{\lambda} e^{-u} du$$

Handwritten notes:  $\lambda t = u$ ,  $\lambda dt = du$

$$= \frac{1}{\lambda} \int_0^{\infty} u e^{-u} du$$

Hence,  $E(T) = \int_{-\infty}^{\infty} t f_T(t) dt = \int_0^{\infty} \lambda t e^{-\lambda t} dt = \frac{1}{\lambda}$

Handwritten notes:  $\lambda > 0$

$$E(T^2) = \int_{-\infty}^{\infty} t^2 f_T(t) dt = \int_0^{\infty} t^2 \lambda e^{-\lambda t} dt = \frac{2}{\lambda^2}$$

Handwritten notes:  $\int_0^{\infty} u^2 e^{-u} du = 2$

$$\text{Var}(T) = E(T^2) - (E(T))^2 = \frac{2}{\lambda^2} - \frac{1}{\lambda^2} = \frac{1}{\lambda^2}$$

Handwritten notes:  $\int_0^{\infty} u^2 e^{-u} du = 2$

The service time  $T$  for a customer follows the exponential distribution with parameter lambda. So, here  $E(t)$  is the expected time of  $t$  expected time of  $t$  is  $-\infty$  to  $\infty$  as we have written as  $t$  times  $f_T(t) dt$  okay this integral  $0$  to  $\infty$  because  $f_T(t)$  the density function  $f_T(t)$  it assumes value  $0$  when  $t$  is  $< 0$ . So, this is  $0$  to  $\infty$  okay  $t$  times lambda  $e$  to the power  $-\lambda t$  dt. Okay now take lambda  $t = u$  then lambda  $dt = du$  okay.

So, what do we get there and the limits because lambda is  $> 0$  as lambda is  $> 0$  okay when  $t$  goes to  $\infty$   $u$  goes to  $\infty$  and when  $t$  is  $0$   $u$  is  $0$  because the limits do not change okay? For  $t$

we put  $u/\lambda$  okay and  $\lambda dt$  we replace by  $du$  so we have  $e$  to the power  $-u$  okay now this is  $1/\lambda$ , so we have  $0$  to infinity  $u e$  to the power  $-u$  okay? Now we can integrate by parts, so we get  $1/\lambda u$  times  $-e$  to the power  $-u$   $0$  to infinity.

And - derivative of  $u$  is  $1$  then  $-e$  to the power  $-u$  okay this will be now  $u$  goes to infinity okay  $u$  times  $e$  to the power  $-u$  goes to  $0$  okay so we have this  $0$  and then when we put  $0$  here for  $u$ , we get this  $0$  okay so what do we get ultimately  $1/\lambda$  and  $-$  becomes  $+$  so  $0$  to infinity  $e$  to the power  $-u$  and this is  $1/\lambda$  we have  $e$  to the power  $-u$  sign okay?  $0$  to infinity so when  $u$  goes to infinity  $t$  is  $0$  when  $u=0$  it is .

So. with  $-$  and  $-$  we get because of power limits we will get it as  $1/\lambda$  and we get  $1/\lambda$  as the mean of the exponential distribution okay in order to determine variants of  $T$  we need to determine  $ET^2$   $ET^2$  is  $\int_{-\infty}^{\infty} T^2 FT dt$  and then this we can similarly determine we have determined  $ET$  okay so this will be  $FT$  is again  $0$  when  $T > 0$  so we can out it as  $\int_0^{\infty} t^2 FT dt$  will be  $\lambda e^{-\lambda t}$  when  $T$  is positive.

Okay we can put again  $t=u$  this will be = putting  $\lambda t = u$  this will be  $= \int_0^{\infty} t^2 \lambda e^{-\lambda t} dt$  will be  $du$  so  $e$  to the power  $-u$  where  $\lambda t = u$  okay this will be  $1/\lambda^2 \int_0^{\infty} u^2 e^{-u} du$  . so, because of  $u^2$  here you have to integrate it twice with respect to  $u$  by integrating by parts and you will see that its value comes out to be  $2$  okay  $2/\lambda^2$ .

We get okay we get  $2/\lambda^2$  variants of  $t$  then  $ET^2 - ET^2$  the whole square okay that is  $2/\lambda^2$   $ET^2$  means  $1/\lambda^2$ , so we get the variants as  $1/\lambda^2$  okay . So, variants is  $1/\lambda^2$  and mean is  $1/\lambda$  okay so mean is  $1/\lambda$  we will use in the examples.

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### Memory less property

The exponential distribution has a characteristic property called the lack of memory property. Suppose  $X$  denotes the length of life of an electronic fuse. Further suppose it is known that fuse has already lasted  $t_0$  units of time, what is the probability that it will fail before  $t$  units,  $t > t_0$ . Then the required probability

$$\begin{aligned}
 P(A|B) &= \frac{P(A \cap B)}{P(B)} \\
 P(\underline{X \leq t} | \underline{X \geq t_0}) &= \frac{P(t_0 \leq X \leq t)}{P(X \geq t_0)} = \frac{P(t_0 \leq X \leq t)}{P(X \geq t_0)} \\
 &= \frac{\int_{t_0}^t \lambda e^{-\lambda x} dx}{\int_{t_0}^{\infty} \lambda e^{-\lambda x} dx} = \frac{(-e^{-\lambda x})_{t_0}^t}{(-e^{-\lambda x})_{t_0}^{\infty}} \\
 &= \frac{e^{-\lambda t_0} - e^{-\lambda t}}{e^{-\lambda t_0}} \\
 &= 1 - e^{-\lambda(t-t_0)}
 \end{aligned}$$

Now let us discuss memory less property the memory the exponential distribution has a characteristic property okay there is a characteristic property of the exponential distribution it is called lack of memory property okay? Suppose  $X$  denotes the length of life of an electronic fuse okay? Length of life often electronic fuse further suppose that it is known that fuse has already lasted  $t_0$  units of time.

What is the probability that it will fail before  $t$  units  $t > t_0$  then the required probability is  $P(X \leq t | X > t_0)$   $X$  denotes the length of life often electronic fuse we want the probability that a it fails before  $t$  units  $t$  units of time so  $X$  must be  $< \text{ or } = t$  but given that it has already lasted  $t_0$  units of time. So,  $X$  is  $> \text{ or } = t_0$  now this is conditional probability that  $X$  is  $< \text{ or } = t$  given that  $X$  is  $> \text{ or } = t_0$ .

We know that probability of  $A$  given  $B = \text{probability of } A \text{ intersection } B / \text{probability of } B$  okay so using that we will have this is event  $A$  this is event  $b$  okay intersection of  $A$  and  $B$  will be  $t_0 \leq X \leq t$  so we will have the probability that  $t_0 \leq X \leq t / P(X \geq t_0)$ . Okay now so this in the numerator what we will have integral  $t_0$  to  $t$  and then the probability density function of the exponential distribution which is  $\lambda e^{-\lambda x}$ .

Okay  $\lambda e^{-\lambda x} dx$  and the denominator will be we have to find the probability that  $X \geq t_0$  so  $t_0$  is  $> \text{infinity}$   $\lambda e^{-\lambda x} dx$  when we

integrate this what we get  $e^{-\lambda t}$  to the power  $- \lambda x$  we get okay and the limits are  $t_0$  to  $t$  here we have again  $e^{-\lambda t}$  to the power  $- \lambda x$  the limits are  $t_0$  to infinity . Okay, So numerator will be  $e^{-\lambda t_0}$   $e^{-\lambda t}$  and denominator.

Because  $\lambda$  is positive  $X$  when goes to infinity it will go to 0  $e^{-\lambda x}$  will go to 0. So, it will be  $e^{-\lambda t_0}$  so we get this okay and this we can write as  $1 - e^{-\lambda(t-t_0)}$ . Okay now you can see the probability that the length of life of the electronic fuse will be at most  $t$  units okay given that it has already lasted  $t_0$  units of time is depends on the length of time  $t - t_0$ . Okay it depends on the length  $t - t_0$  okay.

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Memory less property cont...

Hence this property depends only on the length  $(t - t_0)$  and not on the location of  $t_0$  i.e. the probability that the electronic fuse will fail in the next five minutes given that it has lasted 10,000 hours is the same as the probability that the electronic fuse will fail in the next five minutes given that it has lasted only one hour.

So, what we have this property depends only on this length  $t - t_0$  and not on the location of  $t_0$ . Hence the probability that the electronic fuse will fail in the next five minutes given that it has already lasted 10000 is the same as the probability that the electronic fuse will fail in the next five minutes given that it has the lasted only 1 hour. Okay so suppose the electronic fuse has already lasted 1 hour.

And you have to find the probability that it will fail in the next five minutes that probability is same as the probability that it has already lasted 10000 hours, and it will fail in the next five minutes. Okay so this is called the memory less property the probability does not depend on what has happened before  $t_0$  units of time with only depends on the length  $t-t_0$  okay this is called



memory less property of the exponential distribution. So, that is all in this lecture i thank you very much for your attention.